

# Kalman Filtering on $SO(3)$

Presenter

전준기

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## Section 1

# Kalman Filter



**Figure 1:** Rudolf Emil Kálmán (May 19, 1930 - July 2, 2016)



## Goal

- Tracking problem
  - Find the state of a dynamical system, given the history of observations of the system
  - Filtering: find the **present** state
  - Smoothing: find **past** states
  - Prediction: find **future** states
- Control problem
  - Try to make a dynamical system into a desired state by applying certain actions, given the history of observations of the system
- States, observations, and actions are in some multi-dimensional continua



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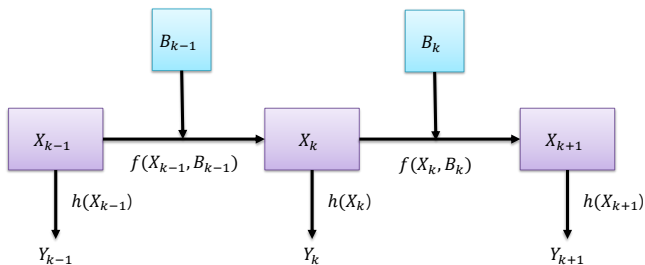
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## Example of System Model



**Figure 2:** Discrete-Time Control System

$X_k$ : system's state at the  $k$ th time instance

$B_k$ : controller's action at the  $k$ th time instance

$Y_k$ : observation of the system at the  $k$ th time instance





## Approaches

- Deterministic approaches
  - “Solve” the equations
  - Minimize corresponding “cost function”
- Probabilistic approaches
  - Parametric vs. Non-parametric
    - Parametric: assumes certain “form” of probability measures
    - Non-parametric: try to find the probability measure itself
  - Bayesian vs. Non-Bayesian
    - Bayesian: parameter itself is a random variable
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- Goal: estimate the state of a system given by a time-series  $(X_k)_{k=0}^{\infty}$ , from observations  $(Y_k)_{k=0}^{\infty}$
- Assumptions:
  - The system evolves linearly:  $X_k = FX_{k-1} + W_k$ 
    - $F$ : a known, fixed linear transformation
    - $W_k$ : **process noise**, driving the system randomly
  - The observation is derived linearly from the state:  $Y_k = HX_k + U_k$ 
    - $H$ : a known, fixed linear transformation
    - $U_k$ : **measurement noise**, making the observation imprecise
  - Further assumptions:
    - $X_0, W_k$ 's,  $U_k$ 's are all independent and all zero-mean Gaussian
    - $W_k$ 's are identically distributed w/ covariance  $Q$
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## Example - Speed Camera

- Want to know: the *speed* of the car
- Observed: the *position* of the car

- $X_k = \begin{bmatrix} P_k \\ V_k \end{bmatrix}$

- $P_k$ : position at  $t = k\Delta t$
- $V_k$ : average velocity between  $t = k\Delta t$  and  $t = (k + 1)\Delta t$

- $\begin{bmatrix} P_k \\ V_k \end{bmatrix} = \begin{bmatrix} 1 & \Delta t \\ 0 & 1 \end{bmatrix} \begin{bmatrix} P_{k-1} \\ V_{k-1} \end{bmatrix} + \begin{bmatrix} 0 \\ A_k \Delta t \end{bmatrix}$

- $A_k$ : random acceleration of the car

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- What is the “best” estimate of  $X_k$  given  $Y_1, \dots, Y_k$ ?
- Since we are doing Bayesian estimate, *average cost* is the concern

### Theorem 1

Given  $L^2$ -r.v.  $X$  and a r.v.  $Y$ , a conditional expectation of  $X$  given  $Y$  is an MMSE (Minimum Mean-Square Error) estimate of  $X$  given  $Y$ ; that is, for any measurable function  $f : \mathcal{Y} \rightarrow \mathcal{X}$ ,

$$\mathbb{E} \left[ \|X - \mathbb{E}[X|Y]\|^2 \right] \leq \mathbb{E} \left[ \|X - f(Y)\|^2 \right]$$

- Suffices to find  $\mathbb{E}[X_k|Y_1, \dots, Y_k]$



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## Kalman Filter (3)

- If  $(X_1, X_2) \sim N \left( \begin{bmatrix} \mu_1 \\ \mu_2 \end{bmatrix} \begin{bmatrix} \Sigma_{11} & \Sigma_{12} \\ \Sigma_{21} & \Sigma_{22} \end{bmatrix} \right)$ , then

$$(X_1 | X_2 = x_2) \sim N \left( \mu_1 + \Sigma_{12} \Sigma_{22}^{-1} (x_2 - \mu_2), \Sigma_{11} - \Sigma_{12} \Sigma_{22}^{-1} \Sigma_{21} \right)$$

- What's the matter?
  - Intractable to directly calculate  $E[X_k | Y_1, \dots, Y_k]$ 
    - Infinite (indefinitely growing) memory requirement
    - Joint distribution of  $(X_k, Y_1, \dots, Y_k)$  is too complicated
- Have to “summarize” the results before taking  $Y_k$  into account
- It turns out,  $\hat{X}_{k-1|k-1} := E[X_{k-1} | Y_1, \dots, Y_{k-1}]$  and  $P_{k-1|k-1} := E[(X_{k-1} - \hat{X}_{k-1|k-1})(X_{k-1} - \hat{X}_{k-1|k-1})^T]$  are sufficient summaries

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  - Intractable to directly calculate  $E[X_k | Y_1, \dots, Y_k]$ 
    - Infinite (indefinitely growing) memory requirement
    - Joint distribution of  $(X_k, Y_1, \dots, Y_k)$  is too complicated
- Have to “summarize” the results before taking  $Y_k$  into account
- It turns out,  $\hat{X}_{k-1|k-1} := E[X_{k-1} | Y_1, \dots, Y_{k-1}]$  and  $P_{k-1|k-1} := E[(X_{k-1} - \hat{X}_{k-1|k-1})(X_{k-1} - \hat{X}_{k-1|k-1})^T]$  are sufficient summaries



## Kalman Filter (4)

- System model:
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    - $Y_k = HX_k + U_k$
  - To compute:  $E[X_k|Y_1, \dots, Y_k]$
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- Computationally cheap; several matrix multiplications and an inversion
- Small memory requirement; only necessary to remember  $\hat{X}_{k|k}$  and  $P_{k|k}$
- Optimal w.r.t. MSE

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- Problems of EKF
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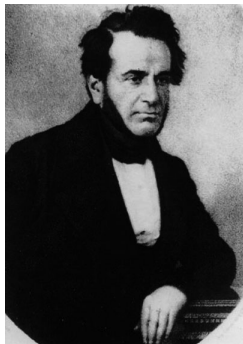


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## Section 2

# Attitude Tracking



**Figure 3:** Olinde Rodrigues (October 6, 1795 - December 17, 1851)

## Problem Definition (1)

- What “attitude” means?
  - The coordinates of three **frame vectors** with respect to the global coordinate system

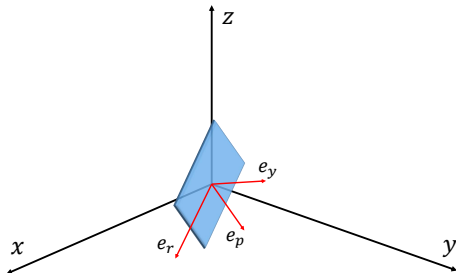
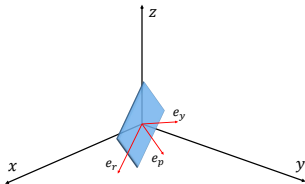


Figure 4: Frame vectors

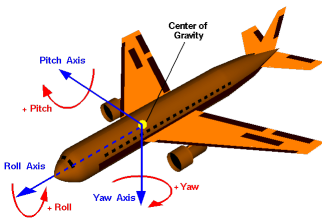
## Problem Definition (2)

- What “attitude” means?
  - (Frame vectors) = **(Rotation matrix)**
  - $r := [e_r \ e_p \ e_y]^T$
  - (Rotation matrix) = **(Orthogonal matrix w/ det.=1)**
- The set of orthogonal matrices w/ det.=1 is called the **special orthogonal group** and denoted as SO(3)
- In summary, we are to find an element in SO(3)



## Problem Definition (3)

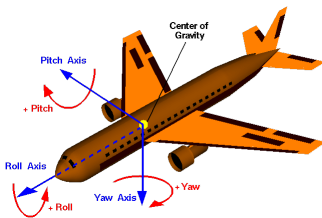
- What we have?
  - **Inertial Measurement Unit (IMU)**: combination of the following three sensors:
    - **Gyroscope**: measures **angular velocity**
    - **Accelerometer**: measures **acceleration**
    - **Magnetometer**: measures **magnetic field**
  - Accelerometer measures the *gravity*
  - Magnetometer measures the *heading*; think of compass



**Figure 5:** The output of gyroscope sensors

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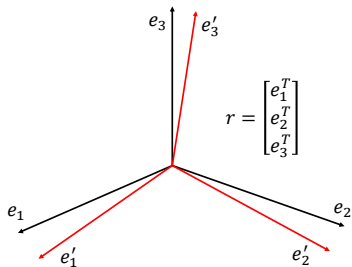


**Figure 5:** The output of gyroscope sensors

## Problem Definition (4)

- Gyroscope measures change of frame vectors *with respect to the local frame vectors*

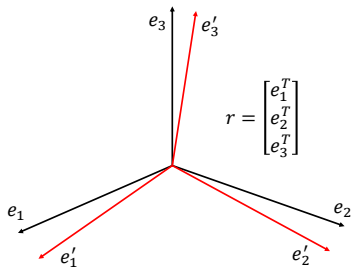
$$\begin{aligned}
 re'_1 &\approx \begin{bmatrix} 1 \\ 0 \\ 0 \end{bmatrix} + \omega_3 \Delta t \begin{bmatrix} 0 \\ 1 \\ 0 \end{bmatrix} - \omega_2 \Delta t \begin{bmatrix} 0 \\ 0 \\ 1 \end{bmatrix} \\
 re'_2 &\approx \begin{bmatrix} 0 \\ 1 \\ 0 \end{bmatrix} + \omega_1 \Delta t \begin{bmatrix} 0 \\ 0 \\ 1 \end{bmatrix} - \omega_3 \Delta t \begin{bmatrix} 1 \\ 0 \\ 0 \end{bmatrix} \\
 re'_3 &\approx \begin{bmatrix} 0 \\ 0 \\ 1 \end{bmatrix} + \omega_2 \Delta t \begin{bmatrix} 1 \\ 0 \\ 0 \end{bmatrix} - \omega_1 \Delta t \begin{bmatrix} 0 \\ 1 \\ 0 \end{bmatrix}
 \end{aligned}$$



## Problem Definition (5)

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$$\begin{aligned}
 r [e'_1 \quad e'_2 \quad e'_3] \\
 \approx I + \begin{bmatrix} 0 & -\omega_3 & \omega_2 \\ \omega_3 & 0 & -\omega_1 \\ -\omega_2 & \omega_1 & 0 \end{bmatrix} \Delta t \\
 \approx \exp([\omega]_{\times} \Delta t)
 \end{aligned}$$



- Hence,  $r(r')^T \approx \exp([\omega]_{\times} \Delta t)$ ,  $r' \approx \exp(-[\omega]_{\times} \Delta t) r$



## Problem Definition (6)

- Evolution equations:

$$\text{Attitude } R_k = \exp(-[\Omega_{k-1}]_{\times} \Delta t) R_{k-1}$$

$$\text{Angular velocity } \Omega_k = \Omega_{k-1} + A_k \Delta t$$

where  $A_k$  is a random *process noise*

- Measurement equations:

$$\text{Gyroscope } G_k = \Omega_k + U_k$$

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where  $U_k, V_k, W_k$  are random *measurement noises*,  $\mathbf{a}$  the constant gravity vector, and  $\mathbf{m}$  the constant Earth magnetic field vector

- Valid only when
  - There is no rapid movement
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  - But we can find MMSE estimator *if* we know the **conditional distribution**
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  - Why Gaussian distribution is so nice?
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      - Affine transforms
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      - Etc.
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- What is the most natural generalization of Gaussian measures on Lie groups?

### Theorem 2 (Drift-Diffusion Equation)

The solution  $f : [0, \infty) \times \mathbb{R}^n \rightarrow \mathbb{C}$  to the differential equation

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## Drift-Diffusion Equation (2)

- Gaussian measures on  $\mathbb{R}^n$  can be characterized as
  - Kernels of drift-diffusion equations, or
  - Measures whose Fourier transforms are of the form

$$\hat{\mu}(\xi) = \exp(-2\pi i \langle \xi, b \rangle - 2\pi^2 \langle \xi, A\xi \rangle)$$

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## Some Backgrounds (1)

### Definition 3 (Unitary representation)

A *unitary representation* of a locally compact Hausdorff group  $G$  is a continuous group homomorphism  $\xi : G \rightarrow U(E)$  into the unitary group of a Hilbert space  $E$  endowed with the strong operator topology.

### Definition 4 (Fourier-Stieltjes transform)

For  $\mu \in M(G)$  and a unitary representation  $\xi$  of  $G$ , the *Fourier-Stieltjes transform of  $\mu$  at  $\xi$*  is defined as

$$\hat{\mu}(\xi) := \int \xi(x^{-1}) d\mu(x)$$

## Some Backgrounds (2)

### Definition 5 (Convolution)

For  $\mu, \nu \in M(G)$ , the *convolution* of  $\mu$  and  $\nu$  is defined as

$$\mu * \nu : A \mapsto \int \int \mathbb{1}_A(xy) d\mu(x) d\nu(y).$$

In particular, according to the embedding  $L^1(G) \rightarrow M(G)$  with respect to the *right Haar measure*,

$$f * \mu : x \mapsto \int f(xy^{-1}) d\mu(y)$$

### Proposition 6

For  $\mu, \nu \in M(G)$  and a unitary representation  $\xi$  of  $G$ ,

$$\widehat{\mu * \nu}(\xi) = \hat{\nu}(\xi)\hat{\mu}(\xi).$$

## Some Backgrounds (3)

### Theorem 7 (Gelfand-Raikov)

*A measure  $\mu \in M(G)$  is uniquely determined by values of its Fourier-Stieltjes transform at irreducible unitary representations; that is, if  $\hat{\mu}(\xi) = 0$  for all irreducible unitary representation  $\xi$  of  $G$ , then  $\mu = 0$ .*

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## Drift-Diffusion Equation (3)

### Theorem 8

Let  $G$  be a Lie group and  $D$  be a left-invariant differential operator on  $G$  given as  $D = -m + \frac{1}{2}\Sigma$ , where  $m \in \mathfrak{g}$  and  $\Sigma \in U(\mathfrak{g})$  is a degree 2 symmetric positive-semidefinite element. Then the unique solution to the differential equation

$$\begin{cases} \frac{\partial f(t,x)}{\partial t} = Df(t,x) \\ f(0,x) = f_0(x), \quad f_0 : G \rightarrow \mathbb{C} \end{cases}$$

is given as

$$f(t,x) = \int f_0(xy^{-1}) d\mu_t(y) = (f_0 * \mu_t)(x),$$

where  $\mu_t$  is the unique measure such that  $\hat{\mu}_t : \xi \mapsto \exp(t\xi_*D)$ .

## Drift-Diffusion Equation (4)

- The theorem says:
  - For any  $\xi$ ,  $\exp(t\xi_*D)$  is a well-defined bounded operator on the Hilber space on which  $\xi$  is defined
  - There uniquely exists a probability measure  $\mu_t \in M(G)$  having  $\xi \mapsto \exp(t\xi_*D)$  as its Fourier-Stieltjes transform
  - $(\mu_t)_{t \geq 0}$  is the kernel of the left-invariant drift-diffusion equation

### Definition 9 (Drift-diffusion measure)

The *drift-diffusion measure* associated to  $D$  is the measure  $\mu_1$ , which is the unique measure satisfying  $\hat{\mu}_1 : \xi \mapsto \exp(\xi_*D)$ .





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## Left-Translated Diffusion Distribution on $SO(3)$ (1)

- Let's focus on the case  $G = SO(3)$
- Let us call a measure  $\mu \in M(G)$  a *left-translated diffusion distribution* if  $\hat{\mu}(\xi) = \exp\left(\frac{1}{2}\xi_*\Sigma\right)\xi(x_0^{-1})$ , and denote  $\mu = LD(x_0, \Sigma)$ 
  - $\Sigma$ , a symmetric positive-definite degree 2 element in  $U(\mathfrak{so}(3))$  will play the role of *covariance matrix*
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- The case when  $\Sigma = \sigma^2\mathbf{1}$ , that is, when it is a *Casimir element*,
  - We call  $\mu$  *central* or *isotropic*
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## Left-Translated Diffusion Distribution on SO(3) (2)

- A kind of central limit theorem hold [2][1]
- For *isotropic* distributions, the pdf can be calculated numerically as

$$f(t) = \sum_{l=0}^{\infty} (2l + 1) e^{-\frac{l(l+1)}{2} \sigma^2} \left( \frac{\sin \left( l + \frac{1}{2} \right) t}{\sin \frac{t}{2}} \right)$$

where  $t$  is the distance from the center

- Is  $x_0$  the MSE estimate when the distribution is  $LD(x_0, \Sigma)$ ?

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## Update Using Gyroscope (1)

- System model:

- $R_k = \exp(-[\Omega_{k-1}]_{\times} \Delta t) R_{k-1}$
- $\Omega_k = \Omega_{k-1} + A_k \Delta t$
- $G_k = \Omega_k + U_k$

- Assume

- $R_0 \perp \Omega_0$
- $R_0 \sim \text{LD}(\bar{r}_0, \Sigma_0)$
- $\Omega_0 \sim \text{N}(\bar{\omega}_0, \sigma_{\Omega,0}^2 \mathbf{1})$

- $(\Omega_1 | G_1 = g_1) \sim \text{N}(\bar{\omega}_1, \sigma_{\Omega,1}^2 \mathbf{1})$  where

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- Similarly,  $(\Omega_0 | G_1 = g_1) \sim \mathcal{N}(\tilde{\omega}_0, \tilde{\sigma}_{\Omega,0}^2 \mathbf{1})$  where

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## Bingham Distribution (1)

### Definition 10 (Bingham distribution)

For a symmetric  $4 \times 4$  real matrix  $M$ , the *Bingham distribution* associated to  $M$  is the probability measure on  $SU(2) \subseteq \mathbb{R}^4$  whose pdf is of the form

$$f(q) = \frac{1}{K(M)} \exp(q^T M q).$$

- $f(q) = f(-q)$ , thus the pushforward of this measure by the covering map  $\pi : SU(2) \rightarrow SO(3)$  doubles the pdf
  - Denote this pushforward onto  $SO(3)$  as  $BH(M)$
- $BH(M) = BH(M + \lambda 1)$ , so we may assume the eigenvalues of  $M$  are  $0 \geq \lambda_1 \geq \lambda_2 \geq \lambda_3$



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## Bingham Distribution (2)

- $f(q) = \frac{1}{K(M)} \exp(q^T M q)$
- $M = P^T \begin{bmatrix} 0 & 0 & 0 & 0 \\ 0 & \lambda_1 & 0 & 0 \\ 0 & 0 & \lambda_2 & 0 \\ 0 & 0 & 0 & \lambda_3 \end{bmatrix} P$ , for some  $P \in O(4)$
- There exists  $q_0 := (s_0, v_0) \in \text{SU}(2) \subseteq \mathbb{R}^{1+3}$  and  $Q \in O(3)$  such that

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- Hence,  $f(q) = \frac{1}{K(M)} \exp \left( (q_0^{-1} q)^T \begin{bmatrix} 0 & 0 \\ 0 & -\frac{1}{2} \Sigma^{-1} \end{bmatrix} (q_0^{-1} q) \right)$

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- $\pi(q_0)$ : unique center, at which  $f$  is maximized
- $\Sigma = Q^T \begin{bmatrix} -\frac{1}{2\lambda_1} & 0 & 0 \\ 0 & -\frac{1}{2\lambda_2} & 0 \\ 0 & 0 & -\frac{1}{2\lambda_3} \end{bmatrix} Q$  is positive-definite
- $Q$ : principal directions
- Eigenvalues: how rapidly spreads along each principal direction
- $K(M) := \int_{\text{SU}(2)} \exp(q^T M q) dq$  depends only on eigenvalues
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  - Numerically verified for some isotropic cases [1]



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## Direction Measurement (1)

- Let  $R \sim \text{BH}(M)$ ,  $V = Rh + U$  for some  $h \in \mathbb{R}^3$  and  $U \sim \text{N}(0, \sigma^2 \mathbf{1})$
- What is the conditional distribution  $R|V$ ?

- Bayes' rule:  $f_{R|V}(r|v) = \frac{f_{V|R}(v|r)f_R(r)}{f_V(v)}$
- $(V|R = r) \sim \text{N}(rh, \Sigma)$ , so  $f_{R|V}(r|v) \propto \exp\left(-\frac{\|v - rhq^{-1}\|^2}{2\sigma^2}\right) \exp(q^T M q) = \exp(q^T (M + M_1) q)$  where

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## Direction Measurement (2)

- In general, for  $V_i = Rh_i + U_i$ ,  $i = 1, \dots, n$ , define

$$M_i := -\frac{1}{2\sigma_i^2} \begin{bmatrix} \|v_i - h_i\|^2 & 2(v_i \times h_i)^T \\ 2(v_i \times h_i)^T & \|v_i + h_i\|^2 \mathbf{1} - 2(v_i h_i^T + h_i v_i^T) \end{bmatrix},$$

then

$$(R|V_1 = v_1, \dots, V_n = v_n) \sim \text{BH} \left( M + \sum_{i=1}^n M_i \right)$$

## Update Using Accelerometer and Magnetometer

- System model:
  - $A_k = R_k \mathbf{a} + V_k$
  - $M_k = R_k \mathbf{m} + W_k$
- Update procedure:
  - 1 Approximate the distribution of  $R_1|G_1$  as  $BH(M)$
  - 2 Calculate  $M_1, M_2$  for  $A_1, M_1$
  - 3 Then the distribution of  $R_1|G_1, A_1, M_1$  is approximately  $BH(M + M_1 + M_2)$
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## Algorithm Summary (1)

- Input: initial distribution, noise variances, constant vectors  $\mathbf{a}$ ,  $\mathbf{m}$
- Output:  $\bar{r}$
- Initialize  $\bar{r} \leftarrow \bar{r}_0$ ,  $\Sigma \leftarrow \Sigma_0$ ,  $\bar{\omega} \leftarrow \bar{\omega}_0$ ,  $\sigma_\Omega^2 \leftarrow \sigma_{\Omega,0}^2$
- For each time instance,
  - 1 Get measurements  $(g, a, m)$
  - 2 Update using  $g$ :

$$\bar{r} \leftarrow \exp \left( - \left[ \frac{(\sigma_U^2 + \sigma_A^2 \Delta t^2) \bar{\omega}}{\sigma_U^2 + \sigma_A^2 \Delta t^2 + \sigma_\Omega^2} + \frac{\sigma_\Omega^2 g}{\sigma_U^2 + \sigma_A^2 \Delta t^2 + \sigma_\Omega^2} \right]_\times \Delta t \right) \bar{r}$$

$$\Sigma \leftarrow \Sigma + \frac{(\sigma_U^2 + \sigma_A^2 \Delta t^2) \sigma_\Omega^2 \Delta t^2}{\sigma_U^2 + \sigma_A^2 \Delta t^2 + \sigma_\Omega^2} \mathbf{1}$$

$$\bar{\omega} \leftarrow \frac{\sigma_U^2 \bar{\omega}}{\sigma_U^2 + \sigma_A^2 \Delta t^2 + \sigma_\Omega^2} + \frac{(\sigma_A^2 \Delta t^2 + \sigma_\Omega^2) g}{\sigma_U^2 + \sigma_A^2 \Delta t^2 + \sigma_\Omega^2}$$

$$\sigma_\Omega^2 \leftarrow \frac{\sigma_U^2 (\sigma_A^2 \Delta t^2 + \sigma_\Omega^2)}{\sigma_U^2 + \sigma_A^2 \Delta t^2 + \sigma_\Omega^2}$$

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- For each time instance,
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- Let  $M_1 := \begin{bmatrix} -\frac{\|a-\mathbf{a}\|^2}{2\sigma_V^2} & -\frac{\|m-\mathbf{m}\|^2}{2\sigma_W^2} & \left(\frac{\mathbf{a}\times a}{\sigma_V^2} + \frac{\mathbf{m}\times m}{\sigma_W^2}\right)^T \\ \frac{\mathbf{a}\times a}{\sigma_V^2} + \frac{\mathbf{m}\times m}{\sigma_W^2} & \frac{\mathbf{a}\mathbf{a}^T + \mathbf{a}\mathbf{a}^T}{\sigma_V^2} + \frac{\mathbf{m}\mathbf{m}^T + \mathbf{m}\mathbf{m}^T}{\sigma_W^2} & -\left(\frac{\|a+\mathbf{a}\|^2}{2\sigma_V^2} + \frac{\|m+\mathbf{m}\|^2}{2\sigma_W^2}\right) \mathbf{1} \end{bmatrix}$
- Find  $q_0 \in \text{SU}(2)$  with  $\pi(q_0) = \bar{r}$
- Let  $M_2 := Q_L(q_0) \begin{bmatrix} 0 & 0 \\ 0 & -\frac{1}{2}\Sigma^{-1} \end{bmatrix} Q_L(q_0^{-1})$
- Find  $(\bar{r}, \Sigma)$  such that  $\text{LD}(\bar{r}, \Sigma) \approx \text{BH}(M_1 + M_2)$

## Discussions

- Why  $(a, m)$  are used only for updating  $R$  but not  $\Omega$ ?
- Incorrect independence assumption  $R_k \perp \Omega_k$
- Consideration of gyroscope bias and accel./magnet. disturbance
- Do claims really hold?
- Extension to  $SE(3)$ ?

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Thank you.

## References



S. Matthies, J. Muller, and G. W. Vinel.

On the normal distribution in the orientation space.

*Textures & Microstructures*, 10:77–96, 1988.



D. I. Nikolayev and T. I. Savyolova.

Normal distribution on the rotation group  $SO(3)$ .

*Textures & Microstructures*, 29:201–233, 1997.

## Derivation of Kalman Filter (1)

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  - $X_k = FX_{k-1} + W_k$
  - $Y_k = HX_k + U_k$
- To compute:  $E[X_k | Y_1, \dots, Y_k]$

- Denote  $\hat{X}_{k|k} := E[X_k | Y_1^k]$
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$$\begin{aligned}\tilde{Y}_k &:= Y_k - E[Y_k | Y_1^{k-1}] \\ &= Y_k - HE[X_k | Y_1^{k-1}]\end{aligned}$$

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- One can show that if  $(X, Y, Z)$  are jointly Gaussian and  $X \perp Y$  then

$$E[Z | X, Y] = E[Z | X] + E[Z | Y] - E[Z],$$

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- To compute:  $E[X_k | Y_1, \dots, Y_k]$
- $\hat{X}_{k|k-1} = E[X_k | Y_1^{k-1}] = FE[X_{k-1} | Y_1^{k-1}] = F\hat{X}_{k-1|k-1}$
- Need to know  $E[X_k | \tilde{Y}_k]$
- According to a well-known formula for jointly Gaussian r.v.s:

$$E[X_k | \tilde{Y}_k] = E[X_k \tilde{Y}_k^T] E[\tilde{Y}_k \tilde{Y}_k^T]^{-1} \tilde{Y}_k$$

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- To compute:  $E[X_k | Y_1, \dots, Y_k]$
- $\tilde{Y}_k = Y_k - E[Y_k | Y_1^{k-1}] = H(X_k - \hat{X}_{k|k-1}) + U_k$ , so
- $E[\tilde{Y}_k \tilde{Y}_k^T] = HP_{k|k-1}H^T + R$ , where
  - $P_{k|k-1} := E[(X_k - \hat{X}_{k|k-1})(X_k - \hat{X}_{k|k-1})^T]$ 
    - This is the **prediction uncertainty**
- Again from independence,  $E[X_k \tilde{Y}_k^T] = P_{k|k-1}H^T$

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## Derivation of Kalman Filter (5)

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- To compute:  $E[X_k | Y_1, \dots, Y_k]$

- Hence,

$$E[X_k | \tilde{Y}_k] = K_k \tilde{Y}_k$$

where

- $K_k := P_{k|k-1} H^T (H P_{k|k-1} H^T + R)^{-1}$ 
  - This is called the **Kalman gain**
- We are to find the prediction uncertainty  $P_{k|k-1} := E[(X_k - \hat{X}_{k|k-1})(X_k - \hat{X}_{k|k-1})^T]$  from now on

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- To compute:  $E[X_k | Y_1, \dots, Y_k]$
- $X_k - \hat{X}_{k|k-1} = F(X_{k-1} - \hat{X}_{k-1|k-1}) + W_k$ , so

$$P_{k|k-1} = FP_{k-1|k-1}F^T + Q$$

where

- $P_{k-1|k-1} := E[(X_{k-1} - \hat{X}_{k-1|k-1})(X_{k-1} - \hat{X}_{k-1|k-1})^T]$
- This is the **estimation uncertainty**
- From  $\hat{X}_{k|k} = \hat{X}_{k|k-1} + E[X_k | \tilde{Y}_k]$ , one can derive the formula

$$P_{k|k} = (I - K_k H)P_{k|k-1}$$

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## Derivation of Kalman Filter (7)

- System model:
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  - $Y_k = HX_k + U_k$
- To compute:  $E[X_k|Y_1, \dots, Y_k]$
  
- In summary, the overall procedure is done in two phases:
  - ① **Prediction phase:** compute  $\hat{X}_{k|k-1}$ ,  $P_{k|k-1}$  from  $\hat{X}_{k-1|k-1}$ ,  $P_{k-1|k-1}$  using the formulas
    - $\hat{X}_{k|k-1} = F\hat{X}_{k-1|k-1}$
    - $P_{k|k-1} = FP_{k-1|k-1}F^T + Q$
  - ② **Update phase:** compute  $\hat{X}_{k|k}$ ,  $P_{k|k}$  from  $\hat{X}_{k|k-1}$ ,  $P_{k|k-1}$ , and  $Y_k$  using the formulas
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